Does foreign sector help forecast domestic variables in DSGE models?

Marcin Kolasa (Warsaw School of Economics, Narodowy Bank Polski) Michał Rubaszek (Warsaw School of Economics)

Motivation

Direct motivation

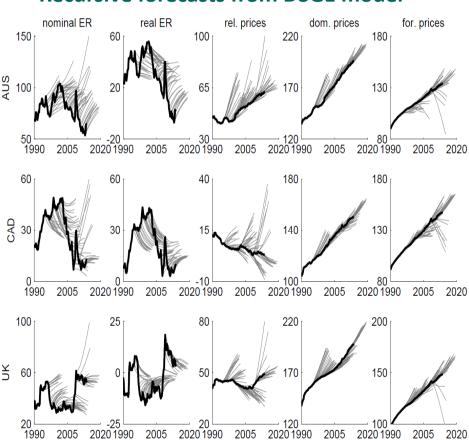
Trigger:

One of the findings in the project "Exchange rate forecasting with DSGE models" is that the inability of open economy DSGE model to forecast NER is due to its failure to forecast domestic and foreign prices

Question:

Does adding the foreign sector to NK model help in forecasting the domestic economy at all?

Recursive forecasts from DSGE model



Source: Ca' Zorzi, Kolasa, Rubaszek (2016)

A quick look at the literature

- > DSGE models: workhorse framework in policy analyses and forecasting
- DSGE model-based forecasts competitive:
 - with time series models (e.g. Smets and Wouters 2007)
 - professional forecasters (e.g. Kolasa, Rubaszek and Skrzypczynski 2012)
 - survey in Handbook on Econ. For. chapter (Del Negro and Schorfheide 2013)
- Vast majority of studies evaluating forecasting performance of DSGE models focus on the US economy and assume a closed economy set-up
- This is at odds with models used in central banks / financial institutions, which are rather open economy models.



A quick look at the literature

[All?] studies that evaluate out-of sample performance of NOEM DSGE models:

- ✓ Adolfson et al. (2007): EA; DSGE vs BVAR; evaluation sample 1994-2002
- ✓ Lees et al. (2007): NZL; DSGE vs BVAR vs RBNZ; evaluation sample 1998-2003
- ✓ Adolfson et al. (2008): Sweden; DSGE vs BVAR; evaluation sample 1999-2004
- ✓ Christoffel et al. (2010): EA; DSGE vs BVAR; evaluation sample 1999-2006
- ✓ Gupta and Kabundi (2010): South Africa; DSGE vs BVAR; evaluation sample 2003-2006
- ✓ Alpanda et al. (2011): South Africa; DSGE vs BVAR; evaluation sample 2003-2009
- ✓ Marcellino and Rychalovska (2014): LUX vs EA; DSGE vs BVAR, 2006-2011

The above studies:

- ✓ compare DSGE to time-seriers models (usually BVAR)
- ✓ evaluate forecasts for an individual country and short sample span
- ✓ are silent about how much we gain by accounting for the external block in DSGE model



Motivation

Our question:

Is it important for forecasts to have external sector in DSGE models

Why important:

- 1. If the only target is to produce accurate forecasts, the use of NOEM model might not be cost efficient unless it produces more accurate forecasts
- 2. There are reasons to be skeptical about the empirical success of the NOEM framework: Justiniano and Preston (2010) demonstrate that it fails to account for the substantial influence of foreign shocks that is identified in VAR studies
- 3. NOEM models have difficulty in explaining swings in exchange rates and current account balances (Engel, 2014; Gourinchas and Rey, 2014), hence might distort the indirect impact of foreign variables on the domestic economy

What we do

- > Start from a standard three-equation **New Keynesian (NK)** closed-economy DSGE setup (with standard features such as habits, sticky prices, etc.)
- Consider two extensions:
 - 1. A small scale open economy model that includes exchange rate and the current account in the set of observables (Lubik and Schorfheide 2007, LS)
 - 2. A medium scale open economy model that also includes foreign sector variables in the set of observables (Justiniano and Preston 2010, JP)
- Analyze their out-of-sample performance in terms of point and density forecasts, both on the basis of data for AUS, CAD and UK over years 1995-2013, as well as using Monte Carlo experiment

Forecasting competition design

JP+ model

Model: Justiniano and Preston (2010, JAE) extended for CA balance shock

Structure: Households optimizing utility; Firms maximizing profits; Local currency

pricing; UIP with risk premium; CB follows Taylor rule

Rigidities: Habits in consumption, price stickiness with indexation,

Incomplete int. financial markets

Shocks: Productivity, import markups, household preferences,

int. risk premium, CA balance, monetary

+ foreign block shocks (output, inflation and IR)

Estimation: Bayesian setup, priors as in Justiniano and Preston (2010)

Observables: output, inflation, 3M interest rate – all three at home and abroad

real exchange rate, terms of trade, CA balance

Other models

JP model: JP+ model without CA shock so that the model is identical

to the original model of Justiniano and Preston (2010)

LS model: JP model without shocks to preferences, import markups and ER risk premium.

Important! foreign variables are treated as unobservable.

Setup closely resembles model by Lubik and Schorfheide (2007)

NK model: Openness parameter at 0 so that we have 3-equation closed economy setup:

$$y_{t} = \frac{1}{1+h} \mathbb{E}_{t} y_{t+1} + \frac{h}{1+h} y_{t-1} - \frac{1-h}{\sigma(1+h)} (i_{t} - \mathbb{E}_{t} \pi_{t+1} - g_{t} + \mathbb{E}_{t} g_{t+1})$$

$$\pi_{t} = \frac{\beta}{1+\beta \delta_{H}} \mathbb{E}_{t} \pi_{t+1} + \frac{\delta_{H}}{1+\beta \delta_{H}} \pi_{t-1} + \frac{(1-\theta_{H})(1-\beta \theta_{H})}{\theta_{H}(1+\beta \delta_{H})} mc_{t}$$

$$i_{t} = \rho_{i} i_{t-1} + \psi_{\pi} \pi_{t} + \psi_{y} y_{t} + \psi_{\Delta y} (y_{t} - y_{t-1}) + m_{t}$$



Prior for steady-state inflation

A critique of DSGE models in the context of inflation forecasting in Faust and Wright (2013) is that its good ex-post performance can be attributed to a tight prior imposed on steady-state inflation:

part of the advantage of the DSGE model [in inflation forecasting] stems from use of a prior specified in light of the full estimation sample -- a prior that probably was at odds with what most agents actually expected at the time

We address this critique and use uninformative prior, namely the prior distribution for steady-state inflation is uniform



Data and forecast evaluation sample

Three countries: UK, CAN, AUS

Observables: Δgdp and Δgdp^* ; Δcpi and Δcpi^* ; i_{3M} and i_{3M}^* ; $\frac{CA}{GDP}$; Δrer , Δtot

Sample: 1975:1 -2013:4

Forecast evaluation: 1995:1-2013:4 (76 obs. for 1q ahead fcsts, 65 for 12q ahead fcsts)

Forecasting scheme: recursive

Weights for foreign variables based on BIS EER indices:

	Australia	Canada	UK	US	euro area	Japan	Coverage
Australia	•	2.4	8.8	32.5	30.2	26.1	74.3
Canada	0.3		2.5	81.5	9.6	6.1	90.8
UK	1.0	2.0	R#3	18.5	70.9	7.5	91.9

Results



Point forecasts: RMSFE

General picture:

- None of the analyzed NOEM models can consistently beat the NK benchmark
- 2. Whenever the RMSFE ratios are statistically signicant, they usually point at the NK model as the preferred forecasting tool

Table 1: Root Mean Squared Forecast Error (RMSFE) for DSGE models

Horizon	Unit	United Kingdom	dom		Canada			Australia	_
	5.1	II.	TP ₊	5	TD	TP+	5	ΠD	TP+
	CTT	10	1 10	CT	0.1	1 10	TD	10	1 10
					Output	15			
1	1.27*	1.11***	1.13***	1.43**	1.21*	1.08**	1.03	1.07***	1.17***
2	1.23*	1.12***	1.15***	1.24**	1.13	1.02	1.03	1.07**	1.24**
4	1.16**	1.08***	1.14**	1.12	1.13**	0.97	0.97	1.08***	1.25***
9	1.14**	1.07***	1.13***	1.03	1.13***	0.95	0.89	1.09***	1.19***
~	1.13**	1.07***	1.12***	0.99	1.12***	0.94	0.82**	1.08**	1.10***
12	1.11	1.07***	1.09***	0.91	1.10***	0.92**	0.72***	1.08***	0.98
					Prices				
-	1.19**	1.06	1.12***	1.21*	0.91	1.06*	1.04	1.03	1.07*
2	1.28**	1.00	1.20***	1.32*	0.89	1.13**	1.06	1.05	1.09*
4	1.33**	0.93	1.30***	1.56**	0.89	1.26**	1.05	1.06	1.11
9	1.34**	0.90	1.36***	1.82**	0.94	1.41***	1.03	1.08	1.09
8	1.30**	0.89	1.37***	1.94**	0.98	1.54***	1.01	1.09	1.04
12	1.22**	06.0	1.40***	2.22**	1.06	1.82***	1.05	1.11*	86.0
				In	Interest ra	rates	- G		
	1.33*	0.86**	0.91	1.13	0.99	86.0	1.01	1.04	1.15
2	1.30*	0.85***	0.95	1.16	0.95	1.05	0.94	1.09*	1.17
4	1.28**	0.89	1.06	1.21	96.0	1.16***	1.00	1.14**	1.21**
9	1.28**	0.95	1.20*	1.21	0.95	1.22***	1.09	1.18***	1.23***
~	1.26**	0.97	1.29**	1.18	0.92	1.27***	1.20	1.22***	1.21***
12	1.25	0.99	1.45***	1.15	.86*	1.35***	1.37**	1.26***	1.18***
Notes: The fi	The figures in the table represent the	table repr		ratios of	ratios of the RMSFE from	FE from a	4	given model in comparison to	parison to

the NK benchmark so that the values below unity indicate that forecasts from a given NOEM variant are more accurate than from the benchmark. Asterisks ***, ** and * denote, respectively, the 1%, 5% and 10% significance levels of the two-tailed Diebold-Mariano test, where the long-run variance is calculated with the THE IRRIGATION OF THE PROPERTY OF THE INVESTED HOLD A SIVER MINUTED IN COMPANISON OF Newey-West method.



Density forecasts: log predictive scores (LPS)

LPS calculated with the method proposed by Adolfson et al. (2007)

General picture:

- 1. Picture comparable to the RMSFE results
- 2. For multivariate density forecasts (3 domestic vars.), NOEM indistinguishable from, or significantly worse than closed economy benchmark

Table 2: Log Predictive Scores (LPS) for DSGE models

					,				
Horizon		United Kingdom	dom		Canada			Australia	
	ST	JP	JP+	ST	Jb	JP+	ST	JP	JP+
					Output				
	-0.15**	-0.07***	-0.08***	-0.35***	-0.18*	-0.06	0.00	-0.06***	-0.09***
2	-0.14*	-0.06	-0.08***	-0.25**	-0.10	0.00	0.01	-0.06***	-0.12***
4		-0.10	-0.12***	-0.20	-0.11^*	0.02	0.05	-0.06***	-0.19***
9	-0.21*	-0.17	-0.16***	-0.17	-0.16*	0.02	0.09	-0.07***	-0.24**
∞	-0.26**	-0.24	-0.20***	-0.21	-0.21**	0.01	0.13**	-0.08***	-0.28**
12	-0.37***	-0.37***	-0.29***	-0.13	-0.30**	-0.01	0.19^{***}	-0.12***	-0.31**
					Prices				
	-0.13***	-0.24***	-0.03***	-0.13	-0.04	-0.07**	-0.18***	-0.07***	***20.0-
2	-0.07***	-0.05**	-0.01*	-0.17	0.00	-0.10***	-0.16***	-0.05***	-0.09***
4	0.00	0.14**	-0.01	-0.21**	0.08*	-0.13***	-0.15***	-0.04**	-0.11***
9	0.03	0.22***	-0.02	-0.24**	0.13***	-0.14***	-0.17***	-0.04**	-0.11***
<u></u>	0.04	0.26***	-0.04	-0.23**	0.16***	-0.15***	-0.19***	-0.05**	-0.09**
12	0.07*	0.27***	-0.09	-0.21**	0.19^{***}	-0.19***	-0.27***	-0.06**	-0.08**
				Int	Interest rates	tes			
-	-0.13***	***90.0-	-0.03***	*90.0-	-0.04**	-0.02	0.04**	-0.04***	-0.03*
2	-0.13***	-0.04**	-0.02	-0.05	-0.02	-0.03	0.03**	-0.05***	-0.03
4	-0.11*	0.01	-0.02	-0.04	0.03	-0.05	-0.02	-0.06***	***90.0-
9	-0.09	0.03	-0.06	-0.05	0.05	-0.09**	-0.09***	-0.07***	-0.07**
∞	-0.06	0.04*	*60.0-	-0.04	0.08***	-0.11***	-0.15***	-0.08***	***20.0-
12	-0.03	0.05**	-0.15***	-0.02	0.13***	-0.14**	-0.26***	-0.10***	**90.0-
			S 80	Thr	Chree variables	ples	E 55		
,	-0.44**	-0.31***	-0.11***	-0.59***	-0.31**	-0.17***	-0.13***	-0.15***	-0.17***
2	-0.41**	-0.09***	-0.06**	-0.66***	-0.23**	-0.11	-0.13**	-0.12***	-0.24**
4	-0.43	0.03	-0.04	-0.85**	-0.18**	0.01	-0.16**	-0.13***	-0.35***
9	-0.46	-0.01	-0.06	***06.0-	-0.18**	0.10	-0.20**	-0.15***	-0.43***
∞	-0.47	-0.05	-0.08	-0.86**	-0.17**	0.16	-0.24***	-0.16***	-0.47***
12	-0.49	-0.16	-0.15*	-0.50	-0.16*	0.23*	-0.36***	-0.21***	-0.50***
Notes: 7	he figures i	n the table	The figures in the table represent the differences of the LPS from a given model in comparison to the	he differenc	es of the L	PS from a.	given mode	I in compar	rison to the

Notes: The figures in the table represent the differences of the LPS from a given model in comparison to the NK benchmark so that positive values indicate that forecasts from a given NOEM variant are more accurate than from the benchmark. Asterisks ***, ** and * denote, respectively, the 1%, 5% and 10% significance levels of the two-tailed Amisano and Giacomini (2007) test, where the long-run variance is calculated with the Newey-West method.

Discussion

Discussion

Why NOEM models might generate more accurate forecasts?

- Economic reason:
 a richer specification might better describe the DGP
- 2. Econometric reason: the information set used estimation is extended for six vars. (JP+); five vars.(JP) or two vars. (LS)

Why NOEM models might potentially generate less accurate forecasts?

- 1. NOEM models (especially foreign block) are misspeciffied
- Extended models contain a larger number of estimated parameters, which inflates estimation forecast error
- 3. Prior for NOEM is centered on wrong values

Discussion

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BVAR check

Does using additional data compensate for the estimation forecast error?

- ➤ Similar comparison for BVAR models that differ in terms of the number of observables (NK, LS, JP and JP+)
- Models estimated on the same set of data as DSGE models (in levels)
- Standard Normal-Wishart prior, overall tightness at 0.2 and decay at 1

General finding:

Once again, NK BVAR performs best

Table 3: Root Mean Squared Forecast Error for BVAR models

Horizon		United Kingdom	gdom	Ü	Canada		A	Australia	ч
	ST	JP	JP+	Γ S	JP	JP+	FS	JP	JP+
	£.				Output	ut			
	1.04	1.00	1.00	1.08	1.07	1.08	1.14**	1.13***	1.18***
2	1.03	0.97	0.97	1.09	1.11	1.17*	1.27***	1.23***	1.27***
4	1.05	0.95	96.0	1.05	1.14	1.23	1.39***	1.31***	1.40***
9	1.08	96.0	0.96	96.0	1:1	1.13	1.48**	1.33***	1.50***
~	1.11	1.00	86.0	0.92	1.08	1.07	1.51***	1.29**	1.54***
12	1.19**	1.15	1.10	0.91	1.02	1.00	1.46***	1.14	1.41**
					Prices	se			
	0.92	0.83**	0.78**	1.05*	1.03	1.06	1.02	0.99	1.11
2	0.91	0.82**	0.75**	1.09*	1.07	1.12*	1.02	0.97	1.15
4	0.91	0.81	0.70	1.15	1.14	1.20**	0.99	0.94	1.08
9	0.93	0.84**	0.68**	1.22	1.22	1.30**	0.95	06.0	1.07
∞	96.0	0.91	0.72**	1.30	1.32*	1.43**	0.94	06.0	1.12
12	1.03	1.03	0.82*	1.46	1.50*	1.64*	1.00	0.99	1.29
				Ir	Interest	rates			
₹ 70 €	1.14*	1.06	1.07	1.13	1.01	1.05	1.00	0.99	1.12
2	1.15	1.12	1.11	1.18	1.01	1.10	1.05	1.05	1.21
4	1.14	1.20	1.08	1.25	1.06	1.20	1.15	1.15	1.24
9	1.22	1.33*	1.15	1.28	1.15	1.29	1.22	1.25	1.29
∞	1.27	1.42*	1.19	1.28	1.20	1.34	1.27*	1.34^{*}	1.48***
12	1.30*	1.47**	1.20*	1.17	1.17	1.30	1.41**	1.59***	1.92***
Notes: The figures in the table represent	he table r	332	the ratios	of the F	MSFE	from a giv	ren model	the ratios of the RMSFE from a given model in comparison to the	ison to the

BVAR variant are more accurate than from the benchmark. Asterisks ***, ** and * denote, respectively, the 1%, 5% and 10% significance levels of the two-tailed Diebold-Mariano test, where the long-run variance is NK (3-variable) benchmark so that the values below unity indicate that forecasts from a given open economy calculated with the Newey-West method.

Monte Carlo

How much we gain if JP+ is true DGP and we center priors at correct values?

- MC experiment, N=100, T=156 (1975:1-2013:4)
- DGP: JP+ model, prior set to reflect CAD moments;
- Forecast evaluation: as in the empirical study (1995:1-2013:4)
- Comparison of JP+ and NK, using RMSFE statistic

General findings:

Forecast esimation error is large enough to roughly offset the potential gains from a better specication and correct priors in the JP+

Table 7: Relative RMSFE of JP+ versus NK - Monte Carlo experiment

TOTAL TENEDRING TENEDRING TOTAL		COLLOTAT	TITLE COULD ON THE PRINCIPLE	Tro divis	211011	
	H=1	H=2	H=1 H=2 H=4 H=6	9=H	H=8	H=8 H=12
			Ou	Output		
Median value	0.95	0.95	0.95	0.95	0.94	0.95
Fraction of <1	0.98	0.89	0.81	0.83	8.0	0.74
Fraction of signif. <1	0.29	0.3	0.26	0.25	0.27	0.26
			Price	Price level		
Median value	0.99	86.0	0.97	0.97	0.97	0.97
Fraction of <1	0.78	0.79	0.78	0.75	0.7	69.0
Fraction of signif. <1	0.11	0.13	0.12	0.16	0.16	0.18
			Intere	Interest rate		
Median value	0.99	0.97	0.97	0.97	0.98	86.0
Fraction of <1	0.78	0.77	0.75	0.68	0.65	0.63
Fraction of signif. <1	0.1	0.13	0.15	0.15	0.13	0.11
Notes: This table presents the RMSFF statistics of the ID+ model relative to the NK model obtained in a	1 mode	relative	ato the	NK mo	del obta	ined in a

Notes: This table presents the KMSFE statistics of the JP+ model relative to the INK model obtained in a Monte Carlo experiment in which the data are generated from the JP+ model with fixed parameters. The significance of differences in the RMSFEs is evaluated with the Diebold-Mariano test at 5% significance level.

Conclusions

Conclusions

- 1. Adding a foreign sector block to a NK DSGE model does not improve or even deteriorates forecasts of domestic variables for AUS, CAD, UK
- 2. Similar result holds for similar BVAR exercise
- 3. Monte Carlo experiment shows that this is partially due to an increase in estimation error, but misspecification may also play a role
- 4. DSGE models are not used just to generate forecasts, but also for simulations. The presence of the foreign block may be highly desired. Still, awareness of possible consequences of including open economy variables for forecast quality is important